

# Consistent nonparametric Bayesian inference for discretely observed diffusions

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# Outline

- The diffusion model: assumptions, basic properties
- Approaches to nonparametric inference
- Nonparametric Bayesian inference
  
- Bayesian data augmentation
- Bayesian inference based on continuous-time data
- Sampling diffusion bridges
- Simulated data example
  
- Consistency result
- Concrete priors
- Sketch of proof
  
- Wish list

# The model

**Model:** stationary, ergodic SDE

$$dX_t = b(X_t) dt + \sigma(X_t) dW_t,$$

with  $\sigma > 0$  *known*,  $b$  the infinite-dimensional parameter of interest.  
(For simplicity, we take  $\sigma \equiv 1$ .)

**Observations:**

$$X_0, X_\Delta, X_{2\Delta}, \dots, X_{n\Delta},$$

with  $\Delta > 0$  *fixed*.

**Aim:** make inference about  $b$  on basis of these observations.

## Some possible approaches

- For small  $\Delta$ ,

$$X_{i\Delta} - X_{(i-1)\Delta} \approx b(X_{(i-1)\Delta})\Delta + Z_i,$$

with  $Z_i$  independent  $N(0, \Delta)$ . Use favorite **nonparametric regression method**. Problem: **bias**.

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- What about a nonparametric version of the data augmentation approach? [REFS]

## Model assumptions more precise - 1

We assume that  $\sigma \equiv 1$  and  $b : \mathbb{R} \rightarrow \mathbb{R}$  is locally integrable.

The SDE generates a one-dimensional diffusion with **scale function**

$$s_b(x) = \int_0^x \exp \left( -2 \int_0^y b(z) dz \right) dy.$$

We assume that

$$\lim_{x \downarrow -\infty} s_b(x) = -\infty, \quad \lim_{x \uparrow \infty} s_b(x) = \infty.$$

Then the SDE generates a **recurrent** diffusion on  $\mathbb{R}$ .

## Model assumptions more precise - 2

The **speed measure** of the diffusion is given by

$$m_b(dx) = \exp\left(2 \int_0^x b(z) dz\right) dx.$$

We assume that  $m_b(\mathbb{R}) < \infty$ .

This ensures the diffusion is **positive recurrent/ergodic** and that  $\mu_b = m_b/m_b(\mathbb{R})$  is the unique **invariant probability measure**.

We assume that the diffusion is started in  $\mu_b$ , so that it is **stationary**.

## Some basic properties of the model - 1

Define the **transition operators** of the diffusion by

$$P_t^b f(x) = \mathbb{E}_x^b f(X_t).$$

Under our assumptions

- **Feller** property:

$$P_t^b : C_b(\mathbb{R}) \rightarrow C_b(\mathbb{R}),$$

- we have **transition densities**  $p_b$  relative to Lebesgue measure:

$$P_t^b f(x) = \int f(y) p_b(t, x, y) dy.$$

## Some basic properties of the model - 2

In particular, we have that the data

$$X_0, X_\Delta, \dots$$

form a (discrete-time) **stationary, ergodic Markov chain**, with unique **invariant distribution**  $\mu_b$  and **transition densities**  $(x, y) \mapsto p_b(\Delta, x, y)$ .

Remark: the transition densities can typically **not** be expressed explicitly in terms of the coefficients  $b$  and  $\sigma$  of the SDE!

# Nonparametric Bayesian inference - 1

The **likelihood** for the data  $X_0, X_\Delta, \dots, X_{n\Delta}$  is given by

$$L_n(b) = \pi_b(X_0) \prod_{i=1}^n p_b(\Delta, X_{(i-1)\Delta}, X_{i\Delta}),$$

where  $\mu_b(dx) = \pi_b(x) dx$ . (Have  $\pi_b(x) \propto \exp\left(2 \int_0^x b(z) dz\right)$ .)

## Nonparametric Bayesian inference - 2

Bayesian approach: put **prior**  $\Pi$  on a collection of drift functions  $\mathcal{B}$  and compute the **posterior**

$$\Pi(B | X_0, X_\Delta, \dots, X_{n\Delta}) = \frac{\int_B L_n(b) \Pi(db)}{\int_{\mathcal{B}} L_n(b) \Pi(db)},$$

for  $B \subset \mathcal{B}$ .

Base further inference on the posterior.

Complication:  $L_n(b)$  can **not** be computed explicitly!

# Bayesian data augmentation - 1

- Likelihood for discrete data intractable.
- Likelihood for **continuous data** ( $X_t : t \in [0, n\Delta]$ ) explicit.
- View continuous segments between the discrete observations as **missing data**.
- Use **Gibbs sampling scheme** to draw from the posterior.

## Bayesian data augmentation - 2

**Bayesian data augmentation** in general (Tanner and Wong (1987)):

Suppose data  $y$  depend on parameter  $\theta$ .

Likelihood  $p(y | \theta)$  intractable. Artificial “data”  $z$  such that  $p(y, z | \theta)$  is tractable.

Posterior:

$$p(\theta | y) = \int p(\theta | y, z)p(z | y) dz,$$

$$p(z | y) = \int p(z | y, \theta)p(\theta | y) d\theta.$$

## Bayesian data augmentation - 3

Consider the Markov chain:

- start in  $(z_0, \theta_0)$ ,
- draw  $z_1$  from  $p(z | y, \theta_0)$ ,
- draw  $\theta_1$  from  $p(\theta | y, z_1)$
- etc.

This chain has an invariant probability distribution, whose marginals are  $p(z | y)$  and  $p(\theta | y)$ .

**Gibbs sampler**

## Bayesian data augmentation - 4

To draw from the discrete data posterior in our setting:

- Choose some **initial estimator**  $b$ .

Repeat a large number of times:

- Sample the **missing continuous segments** between the observations, given the observations and the current value of  $b$ , yielding a full path  $(x_t : t \in [0, n\Delta])$ .
- Draw new  $b$  from the **continuous data posterior**  $\Pi(\cdot | x_t : t \in [0, n\Delta])$ .

After “burn-in”, this yields (approximately) draws from the discrete data posterior.

# Bayesian inference based on continuous-time data - 1

The **likelihood** for the continuous data  $(X_t : t \in [0, T])$  is given by

$$L_T^c(b) = \pi_b(X_0) \exp \left( \int_0^T b(X_t) dX_t - \frac{1}{2} \int_0^T b^2(X_t) dt \right),$$

and the **continuous data posterior** is given by

$$\Pi(B | X_t : t \in [0, n\Delta]) = \frac{\int_B L_{n\Delta}^c(b) \Pi(db)}{\int_{\mathcal{B}} L_{n\Delta}^c(b) \Pi(db)},$$

for  $B \subset \mathcal{B}$ .

## Bayesian inference based on continuous-time data - 2

Clever choice for  $\Pi$ : **Gaussian** or **conditionally Gaussian**. Yields (partial) **conjugacy**.

Simple example: finite Gaussian series prior

Take for  $\Pi$  the law of  $\sum Z_k h_k$ , for deterministic functions  $h_k$  and i.i.d. standard normals  $Z_k$ . Then under the posterior

$$b \sim \sum c_k h_k,$$

where  $c \sim N((I + \Sigma)^{-1}\mu, (I + \Sigma)^{-1})$ , for

$$\mu_k = \int_0^{n\Delta} h_k(X_t) dX_t, \quad \Sigma_{ij} = \int_0^{n\Delta} h_i(X_t) h_j(X_t) dt.$$

# Sampling diffusion bridges - 1

**Basic problem:** Suppose  $X$  solves the SDE

$$dX_t = b(X_t) dt + dW_t, \quad t \in [0, 1].$$

Sample from the conditional distribution of  $X$  given that  $X_0 = u$ ,  $X_1 = v$ .

Recall **Girsanov**:

$$\frac{d\mathbb{P}_X}{d\mathbb{P}_W}(X) = \exp\left(\int_0^T b(X_t) dX_t - \frac{1}{2} \int_0^T b^2(X_t) dt\right).$$

## Sampling diffusion bridges - 2

**General observation:** if  $X$  and  $W$  are conditioned on the same functional, the resulting conditional distributions remain equivalent, and the Radon-Nikodym derivative remains the same up to a constant.

**In particular:** let  $\mathbb{P}_X^{u,v}$  be the law of the diffusion bridge we are after, let  $\mathbb{P}_W^{u,v}$  be the law of the Brownian bridge that goes from  $u$  to  $v$ . Then

$$\frac{d\mathbb{P}_X^{u,v}}{d\mathbb{P}_W^{u,v}}(X) \propto \exp\left(\int_0^1 b(X_t) dX_t - \frac{1}{2} \int_0^1 b^2(X_t) dt\right).$$

Hence: can do **importance sampling** to get **exact simulations** (Beskos and Roberts (2005), Beskos et al. (2006)).

# Simulation example - 1

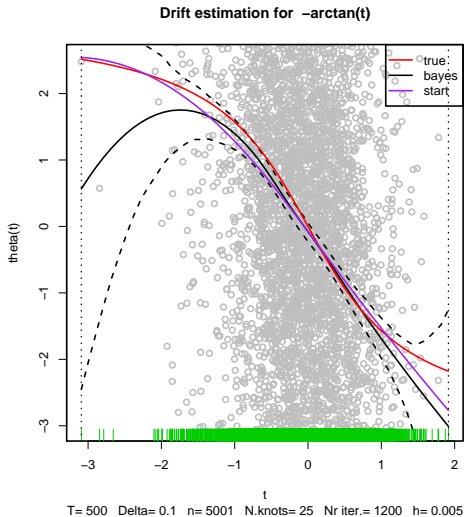


Figure:  $\Delta = 0.1$

## Simulation example - 2

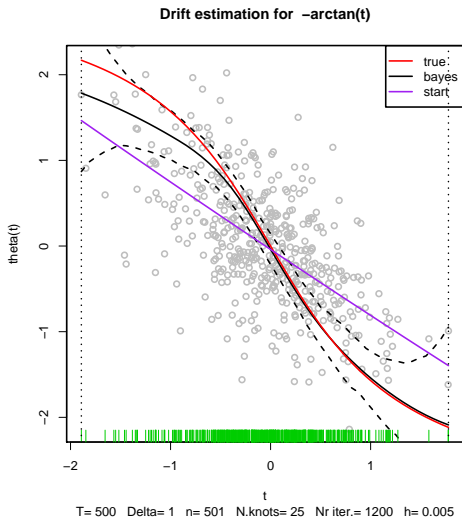


Figure:  $\Delta = 1$

# Consistency - 1

Assume there exists a “true value”  $b_0$ . Does the posterior contract around  $b_0$  as  $n \rightarrow \infty$ ?

General remarks:

- For **nonparametric** models, intuitively reasonable priors may lead to inconsistent procedures (Freedman, Diaconis).
- Existing consistency results for Markov chain data (Tang and Ghosal (2007), Ghosal and Van der Vaart (2007)) are **not applicable**.
- More generally: usual “testing approach” **can not be followed** (yet).

We prove consistency for a weak topology, avoiding testing arguments.

## Consistency - 2

**Topology** on the class  $\mathcal{B}$  of drift functions:

Let  $\nu$  be a finite measure. For  $b \in \mathcal{B}$ ,  $f \in C_b(\mathbb{R})$  and  $\varepsilon > 0$ , let

$$U_{f,\varepsilon}^b = \{b' \in \mathcal{B} : \|P_{\Delta}^{b'} f - P_{\Delta}^b f\|_{1,\nu} < \varepsilon\}.$$

Topology on  $\mathcal{B}$ : such that the collection of sets

$$\{U_{f,\varepsilon}^b : f \in C_b(\mathbb{R}), \varepsilon > 0\}$$

forms sub-base for the neighborhood system at  $b$ .

(i.e.: any open neighborhood of  $b \in \mathcal{B}$  is a union of finite intersections of the form  $U_{f_1,\varepsilon_1}^b \cap \cdots \cap U_{f_m,\varepsilon_m}^b$ .)

## Consistency - 3

Remarks about the topology:

- If  $\nu$  assigns positive mass to non-empty open intervals, then the topology is **Hausdorff**.
- Let  $A_b f = bf' + f''/2$  for a  $C^2$ -function  $f$ . Then for small  $\Delta$ ,

$$P_{\Delta}^{b_1} f - P_{\Delta}^{b_2} f \approx \Delta(A_{b_1} f - A_{b_2} f) = \Delta(b_1 - b_2)f'.$$

Hence for small  $\Delta$ , the constructed topology is close to the  $L^1(\nu)$ -topology.

## Consistency - 4

### Theorem.

Let  $\Pi$  be a prior on a class  $\mathcal{B}$  that is **locally uniformly equicontinuous** and **uniformly bounded**. Then if

$$\Pi\left(b \in \mathcal{B} : \|b - b_0\|_{2, \mu_0} < \varepsilon\right) > 0 \quad \text{for all } \varepsilon > 0,$$

it holds for every open neighborhood  $U_{b_0}$  of  $b_0$  that

$$\Pi(b \notin U_{b_0} \mid X_0, X_\Delta, \dots, X_{n\Delta}) \rightarrow 0 \quad \mathbb{P}_{b_0}\text{-a.s.}$$

as  $n \rightarrow \infty$ .

## Consistency - 5

In words:

If a prior gives mass 1 to a **locally uniformly equicontinuous** and **uniformly bounded** set of drift functions, then we have consistency for all  $b_0$  in the  $L^2(\mu_0)$ -**support** of the prior.

Remark: The basic version of the exact algorithm for simulating diffusion bridges requires that

$$\sup_{b \in \mathcal{B}} (\|b\|_\infty + \|b'\|_\infty) < \infty.$$

Clearly this is stronger than what is required here for consistency.

## Some concrete priors - 1

**Discrete net prior** for arbitrary  $\mathcal{B}$  satisfying the requirements of the theorem:

Fix probability distributions  $(p_n)$  and  $(q_n)$  on  $\mathbb{N}$  and  $0 < \varepsilon_n \downarrow 0$ . Let  $\mathcal{B}_m = \{b|_{[-m,m]} : b \in \mathcal{B}\}$ . By Arzelà-Ascoli,  $\mathcal{B}_m$  is totally bounded for the uniform norm. For every  $n$  we fix a finite  $\varepsilon_n$ -net  $\mathcal{B}_{m,\varepsilon_n} = \{b_1^{m,n}, \dots, b_{k_{m,n}}^{m,n}\}$  for  $\mathcal{B}_m$ . We extend every function in the net to the whole real line by making it constant on  $(-\infty, -m]$  and  $[m, \infty)$ . Define

$$\Pi = \sum_{m=1}^{\infty} \sum_{n=1}^{\infty} \sum_{k=1}^{k_{m,n}} \frac{p_m q_n}{k_{m,n}} \delta_{b_k^{m,n}}.$$

## Some concrete priors - 2

### General construction:

- Fix distribution  $(\rho_m)$  on  $\mathbb{N}$ . For every  $m$ , let  $\Pi_m$  be a prior on  $C_L^\alpha[-m, m]$  ( $\alpha \in (0, 1)$ ,  $L > 0$ ).
- $\bar{C}_L^\alpha[-m, m]$ : extend functions in  $C_L^\alpha[-m, m]$  by making them constant on  $(-\infty, -m]$  and  $[m, \infty)$ . Let  $\bar{\Pi}_m$  be the obvious extension of  $\Pi_m$ .
- Define

$$\Pi = \sum_{m=1}^{\infty} \rho_m \bar{\Pi}_m.$$

If every  $\Pi_m$  has full support relative to the uniform norm and  $b_0 \in C_L^\alpha$ , then we have consistency.

## Some concrete priors - 3

### Integrated Brownian motion:

Fix  $m, L > 0$ ,  $B$  a Brownian motion with standard normal initial law.

- For  $|t| \leq m$ , define  $Y_t = (B_{t+m})_L$ , where

$$(a)_L = \begin{cases} L & \text{if } a \geq L, \\ a & \text{if } |a| \leq L, \\ -L & \text{if } a \leq -L. \end{cases}$$

- Let  $U$  be uniform on  $[-L, L]$ , independent of  $B$ . Let  $\Pi_m$  be the law of

$$Z_t = \left( U + \int_{-m}^t Y_s ds \right)_L, \quad |t| \leq m.$$

## Some concrete priors - 4

### Wavelets:

For an appropriate wavelet basis  $\psi_{j,k}$ , a Hölder ball of functions on  $[-m, m]$  can be represented as

$$\left\{ \sum_{j=1}^{\infty} \sum_{k=1}^{2^j} b_{j,k} \psi_{j,k} : \|b\|_{\infty, \infty}^{\alpha} \leq L \right\},$$

for some  $L > 0$ , where  $\|b\|_{\infty, \infty}^{\alpha} = \sup_j 2^{\alpha j} 2^{j/2} \max_k |b_{j,k}|$ .

Let

$$\Pi_m \sim \sum_{j=1}^J \sum_{k=1}^{2^j} \delta_j U_{j,k} \psi_{j,k},$$

for  $J$ ,  $U_{j,k}$  independent,  $U_{j,k}$  uniformly distributed variables in  $[-L, L]$ , and  $\delta_j = 2^{-\alpha j} 2^{-j/2}$ .

## Sketch of proof - 1

- Enough to show that  $\Pi(B | X_0, \dots, X_{n\Delta}) \rightarrow 0$ ,  $\mathbb{P}_{b_0}$ -a.s., where

$$B = \{b \in \mathcal{B} : \|P_{\Delta}^b f - P_{\Delta}^{b_0} f\|_{1,\nu} > \varepsilon\},$$

with  $\varepsilon > 0$  and  $f$  continuous  $\|f\|_{\infty} \leq 1$ .

- Equicontinuity assumption on  $\mathcal{B}$  implies  $\{P_{\Delta}^b f : b \in \mathcal{B}\}$  is locally uniformly equicontinuous as well.
- Suffices to show that  $\Pi(B' | X_0, \dots, X_{n\Delta}) \rightarrow 0$ ,  $\mathbb{P}_{b_0}$ -a.s., where

$$B' = \left\{ b \in \mathcal{B} : P_{\Delta}^b f(x) - P_{\Delta}^{b_0} f(x) > \eta \quad \forall x \in I \right\},$$

for  $\eta > 0$  and  $I$  a bounded open interval.

## Sketch of proof - 2

- We have

$$\Pi(B' | X_0, X_\Delta, \dots, X_{n\Delta}) = \frac{\int_{B'} L_n(b) \Pi(db)}{\int_{\mathcal{B}} L_n(b) \Pi(db)},$$

- By ergodicity,

$$\frac{1}{n} \log L_n(b) \xrightarrow{\text{as}} -\text{KL}(b_0, b) \gtrsim -\|b - b_0\|_{2, \mu_0}^2.$$

- Hence, for  $n$  large and some  $c > 0$ , the denominator is bounded from below by

$$e^{-nc\delta} \Pi(b : \|b - b_0\|_{2, \mu_0}^2 < \delta)$$

for every  $\delta > 0$ .

## Sketch of proof - 3

- Square root of numerator of expression for posterior:

$$D_n = \sqrt{\int_{B'} L_n(b) \Pi(db)}.$$

- The process

$$M_n = D_n \left(1 - \frac{\eta^2}{8}\right)^{-\sum_{i=1}^{n-1} \mathbf{1}_{X_i \in I}}$$

is a supermartingale.

- By Doob's convergence theorem  $M_n \xrightarrow{\text{as}} M_\infty$ , by ergodicity  $n^{-1} \sum_{i=1}^{n-1} \mathbf{1}_{X_i \in I} \xrightarrow{\text{as}} \mu_0(I)$ .
- Hence  $D_n \xrightarrow{\text{as}} 0$ , exponentially fast.

# Wish list

- estimating  $\sigma^2$  as well
- consistency in stronger topologies
- contraction rates
- optimal priors
- adaptation/model selection
- diffusions in  $\mathbb{R}^d$