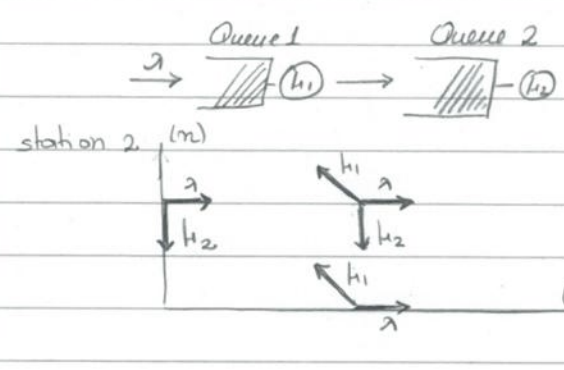


Question: Can we apply the theory of QBD's to processes with infinite phase-space?

Consider a Jackson network consisting of 2-stations and model it as a QBD.



The transition diagram for the tandem network.

Let  $\{(Y_t, J_t), t \geq 0\}$  be the stochastic process modeling the level (number of customers in station 2),  $Y_t$ , and the phase (number of customers in station 1),  $J_t$ , with state space  $S = \{(n, m), n, m \in \mathbb{N}_0\}$ . The corresponding generator matrix  $Q$  assumes the form

$$Q = \begin{bmatrix} \tilde{Q}_1 & Q_0 & & & \\ Q_2 & Q_1 & Q_0 & & \\ & \ddots & \ddots & \ddots & \\ & & & \ddots & \ddots \end{bmatrix}$$

with

$$Q_0 = \begin{bmatrix} 0 & & & & \\ \lambda & 0 & & & \\ & \lambda & 0 & & \\ & & \ddots & \ddots & \ddots \end{bmatrix}, \quad Q_1 = \lambda_2 \mathbb{I}, \quad Q_2 = \begin{bmatrix} -(\lambda + \mu_2) & \lambda & & & \\ & -(\lambda + \mu_1 + \mu_2) & \lambda & & \\ & & \ddots & \ddots & \ddots \\ & & & -(\lambda + \mu_1 + \mu_2) & \lambda \\ & & & & \ddots \end{bmatrix}$$

and  $\tilde{Q}_1 = Q_1 + Q_2$ .

We know that this process is stable if  $\lambda < \min\{\mu_1, \mu_2\}$ . Then, the equilibrium probabilities are  $\pi_{n,m} = (1 - \rho_1) \rho_1^n (1 - \rho_2) \rho_2^m, n, m \geq 0$ .

Consider now that the phase is truncated, at say value  $M$ . We then consider  $\pi_n = (\pi_{n,0}, \pi_{n,1}, \dots, \pi_{n,M})$  for  $n \geq 0$ , and  $\pi_n = \pi_0 R^n, n \geq 0$ , with  $R_M$  the minimal non-negative solution to  $Q_0 + R_M Q_1 + R_M^2 Q_2 = 0$ . For  $M < \infty$ , the decay rate at level  $n$  of the QBD process is  $sp(R_M)$  (spectral radius of  $R_M$ ), i.e.

$$\lim_{n \rightarrow \infty} \frac{\sum_m \pi_{n,m}}{sp(R_M)^n} = c$$

with  $c$  some constant.

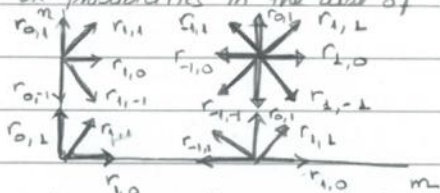
Then, it can be shown, cf. [Corollary 5.5, 1], that, if  $\mu_1 < \mu_2$  then  $sp(R_M)$  increases to  $\eta$  as  $M$  increases, while if  $\mu_1 > \mu_2$  then  $sp(R_M)$  increases to  $\rho_2$  as  $M$  increases. With  $\eta$  the unique solution in  $(0, 1)$  of the function

$$-\lambda - \mu_1 - \mu_2(1 - \eta) + 2\sqrt{\lambda\mu_1} = 0.$$

So, based on this example, we should avoid truncating the state-space of the process. Instead we will investigate other approaches for the calculation of the equilibrium probabilities. Furthermore, we will return to the use of QBD's in a future lecture.

## Generating function approach

Consider a nearest neighbor random walk restricted in the quadrant, with transition rates (OR probabilities in the case of discrete time) as depicted in the figure below



Let  $\pi_{n,m}$  be the equilibrium probabilities, with  $n, m \geq 0$ , and  $P(x,y) = \sum_n \sum_m \pi_{n,m} x^n y^m$ ,  $|x|, |y| \leq 1$ . Multiplying the balance equations with  $x^n$  and  $y^m$ , after summing for  $n, m \geq 0$ , yields a key functional equation for the unknown  $P(x,y)$ , namely

$$K(x,y) P(x,y) = A(x,y) P(x,0) + B(x,y) P(0,y) + C(x,y) P(0,0) \quad (1)$$

where  $K(x,y)$ ,  $A(x,y)$ ,  $B(x,y)$ ,  $C(x,y)$  are known polynomials depending on the parameters of the system. Equations of this type have been studied in the monographs [2] and [3]. The approach suggested consists of mainly two steps:

Step 1: Consider the zeros of the kernel  $K(x,y) = 0$  with  $|x|, |y| \leq 1$ . For such pairs  $P(x,y)$  is analytic, and hence, for these pairs the right hand side of (1) is equal to zero

Step 2: For the pairs  $(x,y)$  satisfying Step 1, one needs to translate the fact that the right hand side of (1) is equal to zero into a Riemann-Hilbert boundary value problem

Unfortunately, the above steps do not constitute a simple, straightforward recipe. In order to illustrate the procedure we consider the  $2 \times 2$  switch.

### Intermezzo

Let  $X$  a discrete r.v. with pdf  $(p_n)$ , then the probability generating function (PGF) is defined as  $P(z) = \sum_{n=0}^{\infty} p_n z^n = E(z^X)$ ,  $|z| \leq 1$ .

Properties

i)  $P(1) = 1$

ii)  $p_n = \frac{P^{(n)}(0)}{n!}$

iii)  $E\left(\frac{x!}{(x-n)!}\right) = P^{(n)}(x)$

iv)  $P(e^z) = M(z) = E(e^{zX})$  which is the moment generating function

v) Setting  $z := z+1$ , yields

$$P(z+1) = E((z+1)^X) = E\left(\sum_{k=0}^{\infty} \binom{X}{k} z^k\right) = \sum_{n=0}^{\infty} p_n \sum_{k=0}^n \binom{n}{k} z^k = \sum_{k=0}^{\infty} z^k \sum_{n=k}^{\infty} p_n \binom{n}{k}$$

$$= \sum_{k=0}^{\infty} z^k m_{(k)}, \text{ with } m_{(k)} = E[X(X-1)\dots(X-k+1)]$$

The latter is known as the factorial moment generating function.

e.g. M/M/1 queue with  $P(z) = \frac{1-\rho}{1-\rho z}$ ,  $|z| < \frac{1}{\rho}$ . Then

$$P(z+1) = \frac{1-\rho}{1-\rho z} = \frac{1}{1-\frac{\rho}{z}} = \sum_{k=0}^{\infty} \left(\frac{\rho}{z}\right)^k z^k$$

$$\Rightarrow m_{k+1} = \left(\frac{\rho}{1-\rho}\right)^k k!, \quad k=0,1,\dots$$

This last equation produces all moments, e.g.  $E(X) = m_{(1)} = \frac{\rho}{1-\rho}$ ,  
 $E(X^2) = E(X(X-1)) + E(X) = m_{(2)} + m_{(1)}$ , etc.

Note that the factorial moment generating function converges in  $|z| < \frac{1}{\rho}$ .

vi) Reconstruction of the distribution from all the moments.

Only, under certain conditions, e.g. if all moments exist and the moment generating function exists in some neighborhood of 0, cf. [page 65, 5]

Counter-examples violating conditions

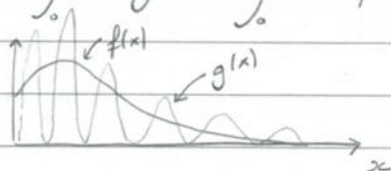
$$X \sim f(x) = \frac{1}{\sqrt{2\pi}x} \exp\left\{-\frac{1}{2} \ln^2(x)\right\}, \quad x > 0$$

$$\text{with } E(X^n) = e^{\frac{1}{2}n^2}, \quad n \in \mathbb{N} \text{ and } E(e^{tx}) = \infty, \quad t > 0.$$

$$Y \sim g(x) = f(x) [1 + \sin(2\pi \ln x)], \quad x > 0$$

$$\text{with } E(Y^n) = e^{\frac{1}{2}n^2}, \quad n \in \mathbb{N} \text{ and } g \text{ is a pdf.}$$

Furthermore,  $\int_0^{\infty} x^n g(x) dx = \int_0^{\infty} x^n f(x) dx$ , however  $g(x) \neq f(x)$



v) Numerical inversion of PGFs, see, eg [4], to find  $p_n$ .

a) by inspection:  $p_n =$  coefficient of  $z^n$  term in  $P(z)$

b) using Taylor expansion around  $z=0$ :  $p_n = \frac{1}{n!} P^{(n)}(0)$

c) using Cauchy's residue theorem (contour integral in the complex plane):

$$p_n = \frac{1}{2\pi i} \oint \frac{P(z)}{z^{n+1}} dz$$

d) using partial fraction method:

$$P(z) = \frac{a_m z^m + a_{m-1} z^{m-1} + \dots + a_1 z + a_0}{b_d z^d + b_{d-1} z^{d-1} + \dots + b_1 z + b_0} = \frac{N(z)}{D(z)}$$

$$= \sum_{k=0}^{m-d} c_k z^k + \frac{\tilde{N}(z)}{D(z)}$$

Let  $D(z) = b_d (z-z_1) \dots (z-z_d)$  if  $D(z)$  has  $d$  distinct roots

Then

$$P_n \sim \sum_1 \alpha_i^{n+1}$$



with  $\alpha_i = \frac{1}{z_i} > \frac{1}{z_i} \forall i \neq 1$  and  $f_1 = -\frac{\tilde{N}(z_1)}{D'(z_1)}$ .

If the dominant root  $(z_1)$  has multiplicity 2, then  $p_n \sim f_{1,2} n \alpha_1^{n+2}$   
 with  $f_{1,2} = \frac{(-1)^2 2! \tilde{N}(z_1)}{D^{(2)}(z_1)}$ .

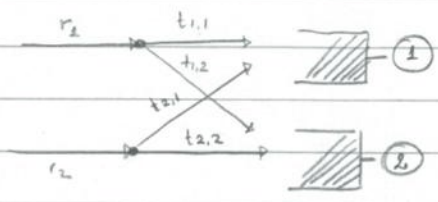
This can be generalized to any multiplicity of the dominant root.

e) Recursion method

$$P(z) = \frac{N(z)}{D(z)} = \frac{\sum_{k=0}^{\infty} \sum_{j=0}^d P_k b_j z^{kj}}{\sum_{j=0}^d b_j z^j} \quad \text{with} \quad \sum_{j=0}^{\min\{d, i\}} P_{i-j} b_j = \begin{cases} a_i, & i \leq n \\ 0, & i > n \end{cases}$$

Then,  $p_n = \frac{1}{b_0} \left[ a_n - \sum_{j=1}^{\min\{d, n\}} b_j p_{n-j} \right], n \geq 0$ .

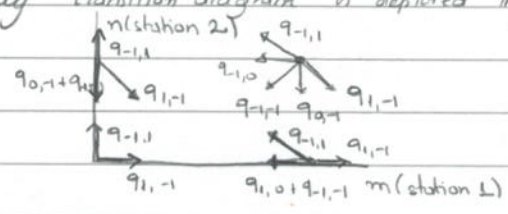
Boundary value method example: 2x2 switch, cf. [6]



Consider the 2x2 clocked buffered switch with 2 input and 2 output ports. Modelling this as a DTMC:

- $r_i$  = prob. to have an arrival of type  $i$ , with  $r_1 = r_2 = p$
- $t_{ij}$  = prob. a job of type  $i$  joins server  $j$ , with  $t_{i,j} = \frac{1}{2}$
- a server serves exactly one job per time unit

The probability transition diagram is depicted in the following figure.



with  $q_{1,-1} = r_1 r_2 + t_{1,1} t_{2,1}$ ,  $q_{0,-1} = r_1 (1-r_2) t_{1,1} + r_2 (1-r_1) t_{2,1}$ ,  $q_{-1,0} = r_1 (1-r_2) t_{1,2} + r_2 (1-r_1) t_{2,2}$   
 $q_{0,0} = r_1 r_2 (t_{1,1} t_{2,2} + t_{1,2} t_{2,1})$ ,  $q_{-1,1} = r_1 r_2 t_{1,2} t_{2,2}$ ,  $q_{-1,-1} = (1-r_1)(1-r_2)$

Let  $P(x,y) = \sum_{m=0}^{\infty} \sum_{n=0}^{\infty} p_{m,n} x^m y^n$ ,  $|x|, |y| \leq 1$  be the PGF and  $p_{m,n}$  the bi-variate equilibrium probability of having  $m$  (respectively  $n$ ) customers in station 1 (2 respectively 2).

From the balance equations, it is straight forward to show that  $P(x,y)$  satisfies the following functional equation

$$(xy - r(x,y)) P(x,y) = (y-1) r(x,0) P(x,0) + (x-1) r(0,y) P(0,y) + (x-1)(y-1) r(0,0) P(0,0), \quad |x|, |y| \leq 1 \quad (2)$$

with  $r(x,y) = (1-p + \frac{p}{2}(x+y))^2$ .

Step 1: Consider the zeros of the kernel  $xy - r(x,y) = 0$ . This gives us a very rich choice of zeros. Because of the symmetry of the model, we choose  $y = \bar{x}$  (conjugate), then the zeros of the kernel reduce to  $|x|^2 - r(x, \bar{x}) = 0$  and let

$$E = \{x \in \mathbb{C} : |x|^2 - r(x, \bar{x}) = 0, |x| \leq 1\}$$

Note that  $E$  forms an ellipse.

Step 2: For  $y = \bar{x}$  and  $x \in E$ , Equation (2) assumes the form

$$0 = (\bar{x}-1) r(x,0) P(x,0) + (x-1) r(0,\bar{x}) P(0,\bar{x}) + (x-1)(\bar{x}-1) r(0,0) P(0,0) \stackrel{x \neq 1}{\Rightarrow}$$

$$0 = \frac{r(x,0) P(x,0)}{x-1} + \frac{r(0,\bar{x}) P(0,\bar{x})}{x-1} + r(0,0) P(0,0), \quad x \in E \setminus \{1\} \quad (3)$$

Let  $g(x) = \frac{r(x,0) P(x,0)}{x-1} + \frac{1}{2} r(0,0) P(0,0)$ , then from (3), together with  $P(x,0) = P(0,\bar{x})$   
 $g(x) + g(\bar{x}) = 0 \Rightarrow 2 \operatorname{Re} g(x) = 0, \quad x \in E \setminus \{1\}$

So, function  $g(x)$  is analytic inside  $E$ , has a simple pole at  $1$ , since

$$\lim_{x \rightarrow 1} (x-1)g(x) = \lim_{x \rightarrow 1} r(x,0)P(x,0) = r(1,0)P(1,0),$$

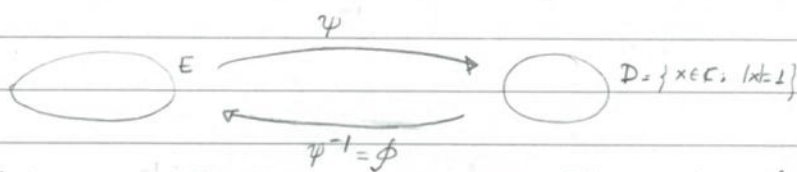
and on the boundary of  $E$  it has zero real part. So the problem reduces in identifying such a function  $g(x)$ .

If we would be looking at the unit circle, say  $D$ , then the problem of identifying the function is trivial; it can be easily verified that

$$h(w) = c \frac{w+1}{w-1}, \quad |w| < 1 \tag{4}$$

has a simple pole at  $1$  and on the boundary  $\text{Re } h(w) = 0$  for  $|w|=1$  and  $w \neq 1$ .

Now consider a function that maps the ellipse into the unit circle, say  $\psi$ ,



that is  $\psi: E \rightarrow D$ , so  $w = \psi(x)$ . Then, (4) yields

$$h(\psi(x)) = c \frac{\psi(x)+1}{\psi(x)-1}$$

if, on top of that, we choose  $\lim_{w \rightarrow 1} (w-1)h(w) = \lim_{x \rightarrow 1} (\psi(x)-1)g(x) \Rightarrow 2c = r(1,0)P(1,0)\psi'(1)$  because of the mapping  $\psi(1)=1$

$$\text{then } g(x) = \frac{r(1,0)P(1,0)}{2} \frac{\psi(x)+1}{\psi(x)-1} \psi'(x) \tag{5}$$

$$\begin{aligned} (2) \stackrel{(5)}{\Rightarrow} (xy - r(x,y))P(x,y) &= (x-1)(y-1) \left[ g(x) + g(y) \right] = \frac{r(1,0)P(1,0)}{2} (x-1)(y-1) \left[ \frac{\psi(x)+1}{\psi(x)-1} + \frac{\psi(y)+1}{\psi(y)-1} \right] \\ &= \frac{r(1,0)P(1,0)}{2} (x-1)(y-1) \frac{\sqrt{2}(\psi(x)\psi(y)-1)}{(\psi(x)-1)(\psi(y)-1)} \psi'(x) \end{aligned}$$

$$\Rightarrow P(x,y) = r(1,0)P(1,0) \frac{(\psi(x)\psi(y)-1)}{xy - r(x,y)} \frac{(x-1)(y-1)}{(\psi(x)-1)(\psi(y)-1)} \psi'(x)$$

Lastly we show that  $r(1,0)P(1,0) = 1-p$ . Note that for  $x=1$ , (2) yields

$$(y - r(1,y))P(1,y) = (y-1)r(1,0)P(1,0) \Rightarrow P(1,y) = r(1,0)P(1,0) \frac{y-1}{y - r(1,y)}$$

$$\begin{aligned} P(1,y) &= r(1,0)P(1,0) \frac{y-1}{y - (1-p+py)} = r(1,0)P(1,0) \frac{1}{\frac{y - (1-p+py)}{y-1}} \\ \stackrel{y=1}{\Rightarrow} 1 &= r(1,0)P(1,0) \frac{1}{\frac{1 - (1-p+p)}{1-1}} = r(1,0)P(1,0) \frac{1}{\frac{-p^2 + 1 - p + p^2}{4}} \Rightarrow r(1,0)P(1,0) = 1-p \end{aligned}$$

$$\text{Furthermore } P(1,y) = \frac{(1-p)}{(1-p/2)^2} \frac{1}{1 - \frac{p^2}{4}y} = \frac{1-p}{(1-p/2)^2} \sum_{n=0}^{\infty} \left( \frac{p/2}{1-p/2} \right)^{2n} y^n$$



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